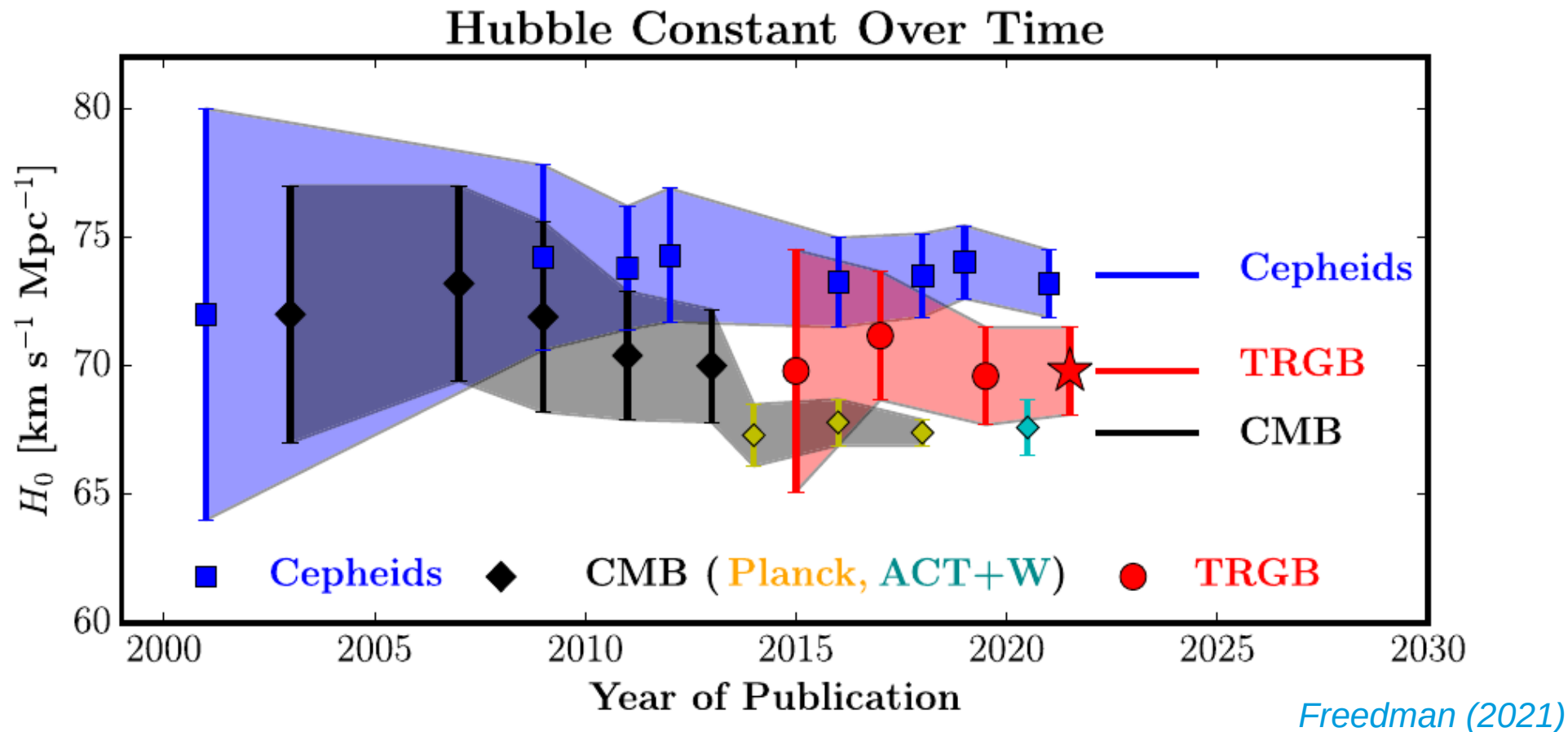


# The Hubble tension survey: A statistical analysis of the 2012-2022 measurements

Bao Wang, Martin Lopez-Corredoira, Jun-Jie Wei  
2024, MNRAS. 527, 7692

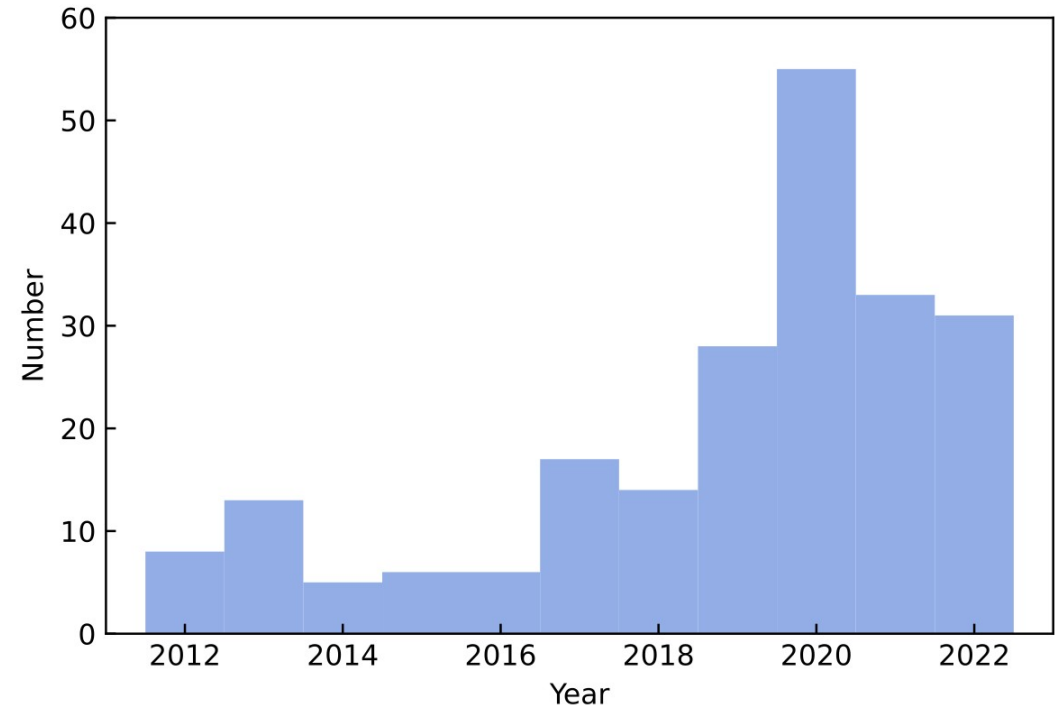
# The Hubble Tension

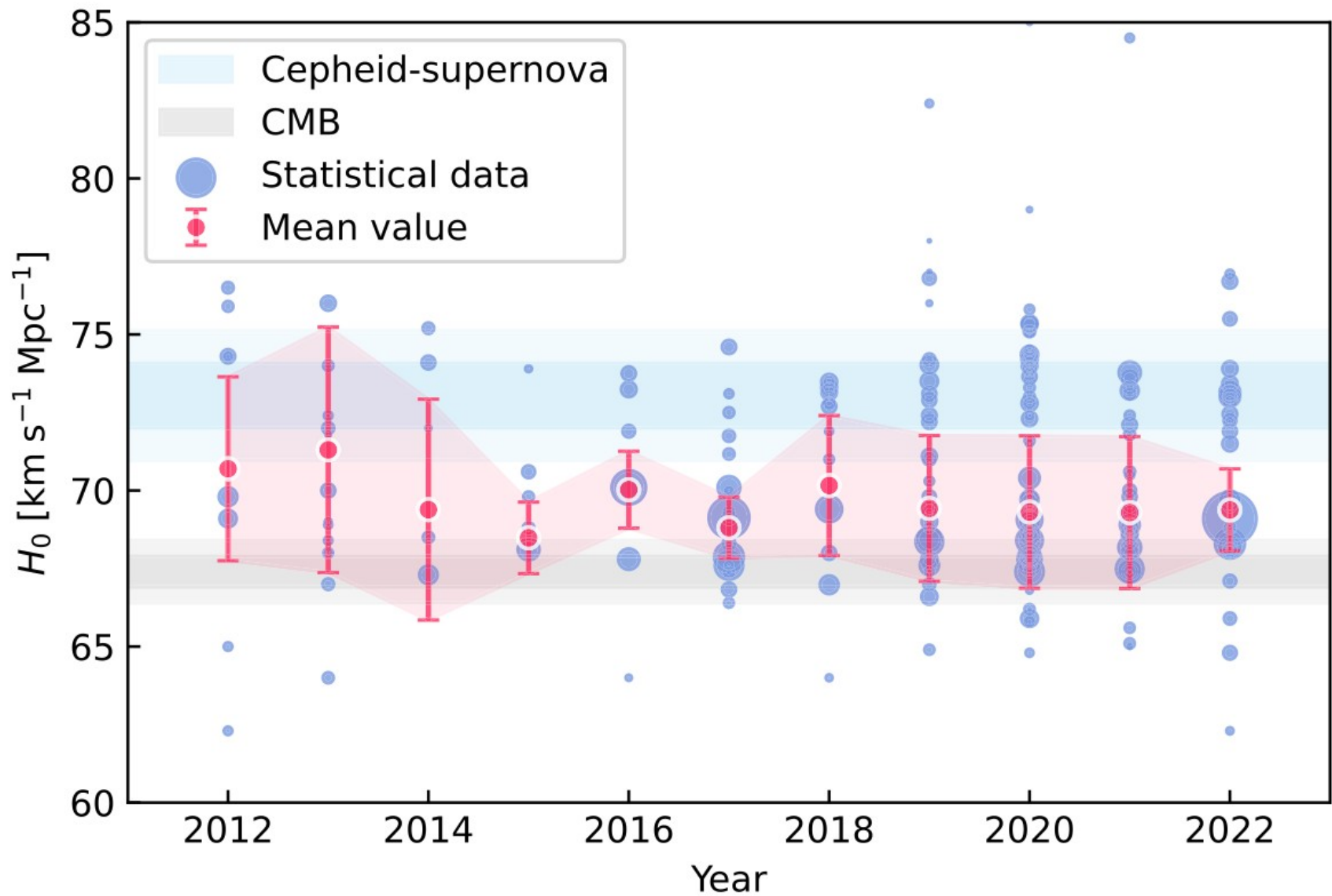


# DATA

## 216 $H_0$ measurements made between 2012 – 2022:

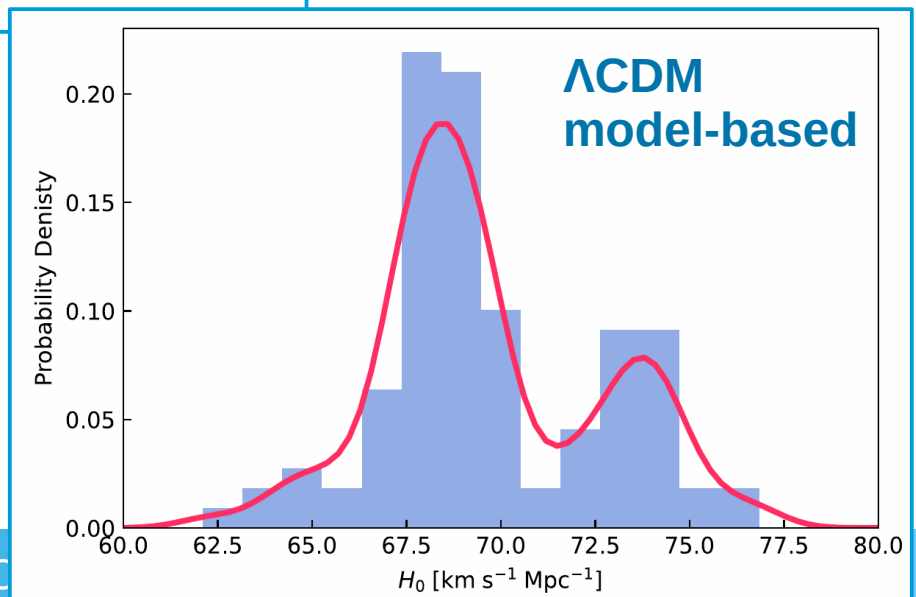
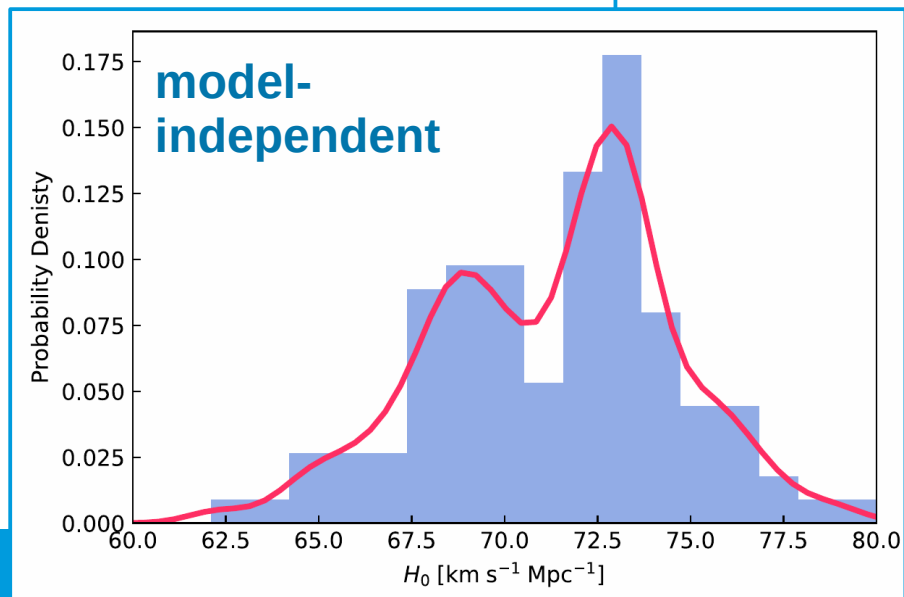
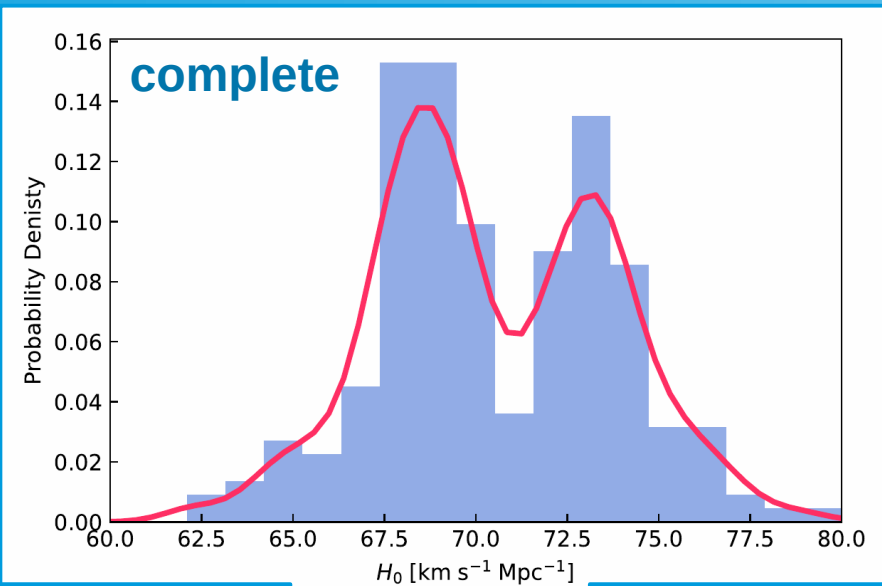
- **109** from model-independent methods (Cepheids + SNe Ia distance ladder)
- **107** from CMB under the **standard**  $\Lambda$ CDM model





Data divided into 3 samples:

- **complete** (216)
- **model-independent** (109)
- **$\Lambda$ CDM model-based** (107)



# 1. Statistical significance of bimodality

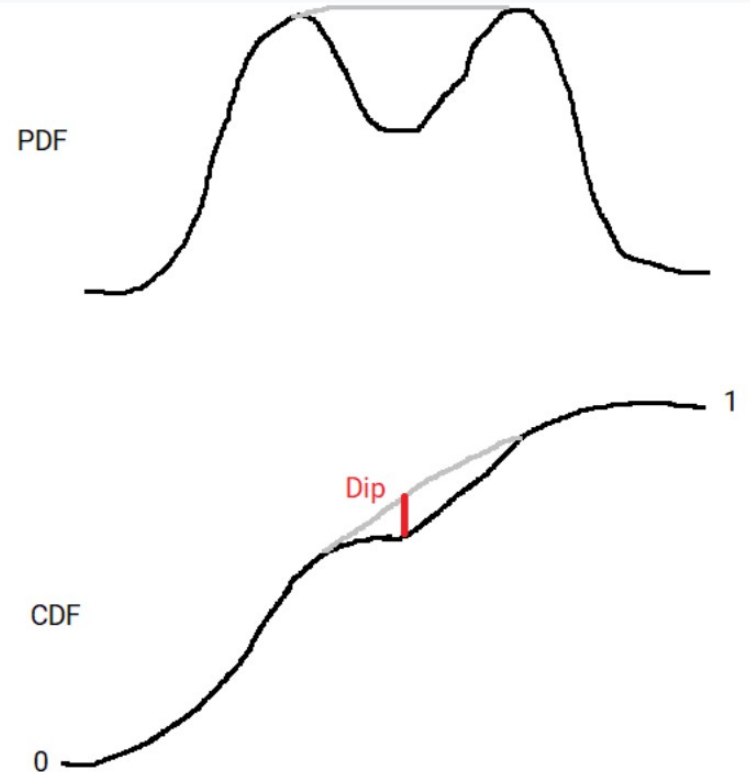
## The „Dip Test”

- used to test multimodality of distributions (Hartigan & Hartigan 1985) by calculating the **discrepancy** between:
  - the **empirical distribution function** and
  - the **unimodal distribution that minimizes the maximum discrepancy**

# 1. Statistical significance of bimodality

## The „Dip Test”

- the distribution can be deformed into a unimodal one by moving the CDF **by at most the dip** at each point
- the **Dip** is the **smallest number** for which this is true



<https://skeptric.com/dip-statistic/index.html>



# 1. Statistical significance of bimodality

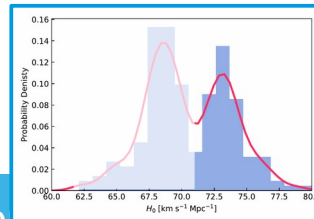
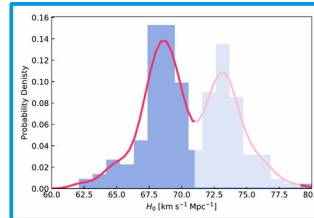
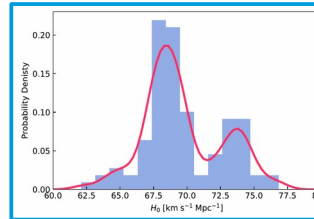
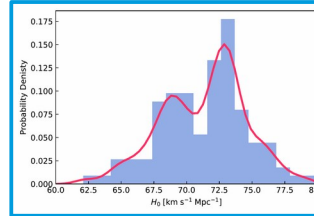
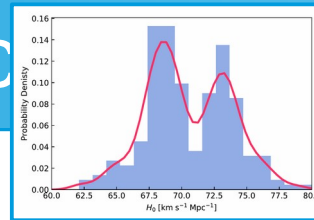
## The „Dip Test”

- the ***p* value** for the dip test – the probability of unimodality:
  - $p < 0.05$  – significant multimodality
  - $0.05 < p < 0.10$  – marginal multimodality
  - $p \gg 0.05$  – unimodality

# 1. Statistical significance of bimodality

Data divided into 5 samples:

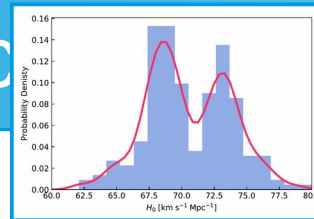
- **complete** (216)
- **model-independent** (109)
- **$\Lambda$ CDM model-based** (107)
- **$H_0 < 71 \text{ km s}^{-1} \text{ Mpc}^{-1}$**  (118)
- **$H_0 \geq 71 \text{ km s}^{-1} \text{ Mpc}^{-1}$**  (98)



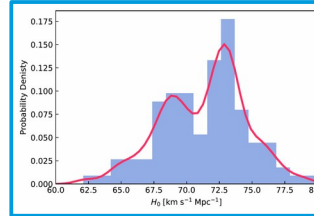
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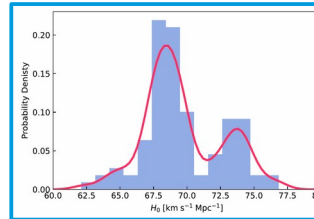
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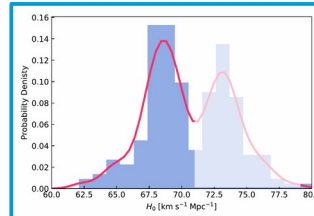
$$p = 0.01$$



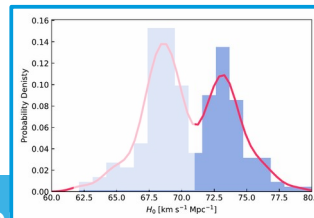
$$p = 0.46$$



$$p = 0.16$$



$$p = 0.96$$



$$p = 0.98$$

## 2. Statistical significance of measurements

## 2. Statistical significance of measurements

$$\chi^2 = \sum_{i=1}^N \frac{(H_{0,i} - \overline{H_0})^2}{\sigma_i^2}$$

calculated for the same 5 subsamples

# 2. Statistical significance of measurements

**Table 1.** The results of dip tests, weighted averages and  $\chi^2$  fittings in three categories: complete, model-independent, and  $\Lambda$ CDM model-based measurements. The  $Q$  values measure the statistical significance for  $\chi^2$  fittings, and the  $p$  values measure the significance of the unimodality for dip tests.

	Number	$\overline{H_0}$ ( $\sigma$ ) ( $\text{km s}^{-1} \text{Mpc}^{-1}$ )	$\chi^2$	$Q$	$p$
Complete	216	$69.35 \pm 0.12$	515.99	$8.85 \times 10^{-27}$	0.01
Model-independent	109	$70.82 \pm 0.22$	181.48	$1.23 \times 10^{-5}$	0.46
$\Lambda$ CDM model-based	107	$68.94 \pm 0.13$	237.56	$4.36 \times 10^{-12}$	0.16
$H_0 < 71 \text{ km s}^{-1} \text{Mpc}^{-1}$	118	$68.78 \pm 0.07$	95.09	0.93	0.96
$H_0 \geq 71 \text{ km s}^{-1} \text{Mpc}^{-1}$	98	$73.53 \pm 0.13$	30.05	$\sim 1.00$	0.98

$Q > 0.05$  means „statistically significant”

$Q$  measures the probability that trends come from chance errors  
 $\implies$  it is unlikely that the trends in 3 groups are due to chance errors

# 3. Outlier rejection

How many outliers should we remove to reach  $Q > 0.05$  ?

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How many outliers should we remove to reach  $Q > 0.05$  ?

define  $x$  – the number of  $\sigma$  deviations between the measurements  $H_{0,i}$  and the average  $\overline{H_0}$

$$x = \frac{|H_{0,i} - \overline{H_0}|}{\sigma_i}$$

then exclude data with  $x > x_{min}$  and repeat the analysis



# 3. Outlier rejection

**Table 2.** The results of weighted averages and  $\chi^2$  fittings in three categories after removing the outliers. The number of outliers and minimal deviations are also displayed.

	$x_{\min}$	Outliers	Number	$\overline{H_0}$ ( $\sigma$ ) ( $\text{km s}^{-1} \text{Mpc}^{-1}$ )	$\chi^2$	$Q$
Complete	2.4	27	189	$69.17 \pm 0.09$	216.42	0.08
Model-independent	3.6	1	108	$72.45 \pm 0.21$	84.20	0.95
$\Lambda$ CDM model-based	2.6	13	94	$68.85 \pm 0.10$	106.33	0.16

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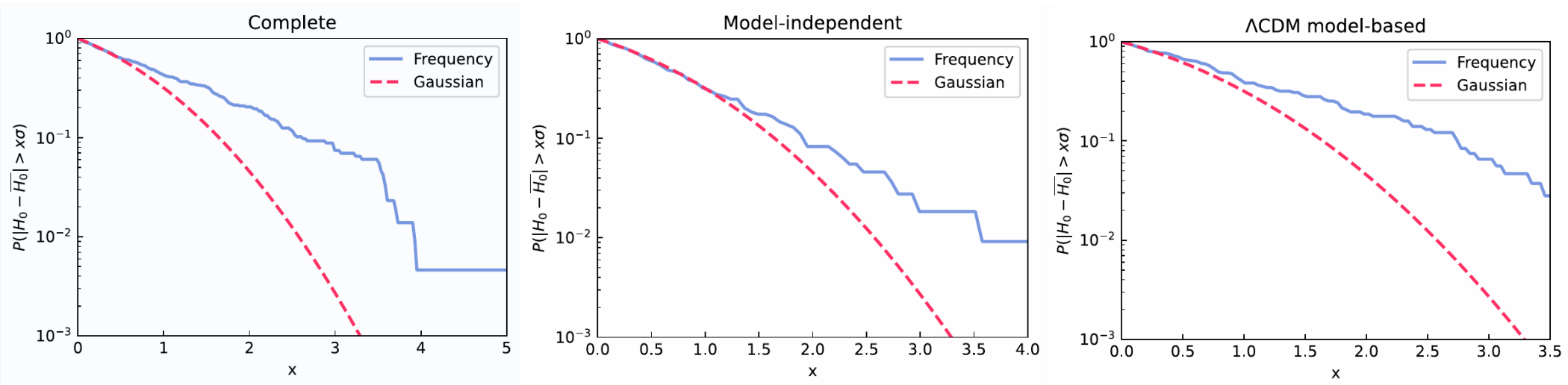
a sign of a possible tension

==> what is the degree of the tension?

# 4. Estimating the real degree of tension

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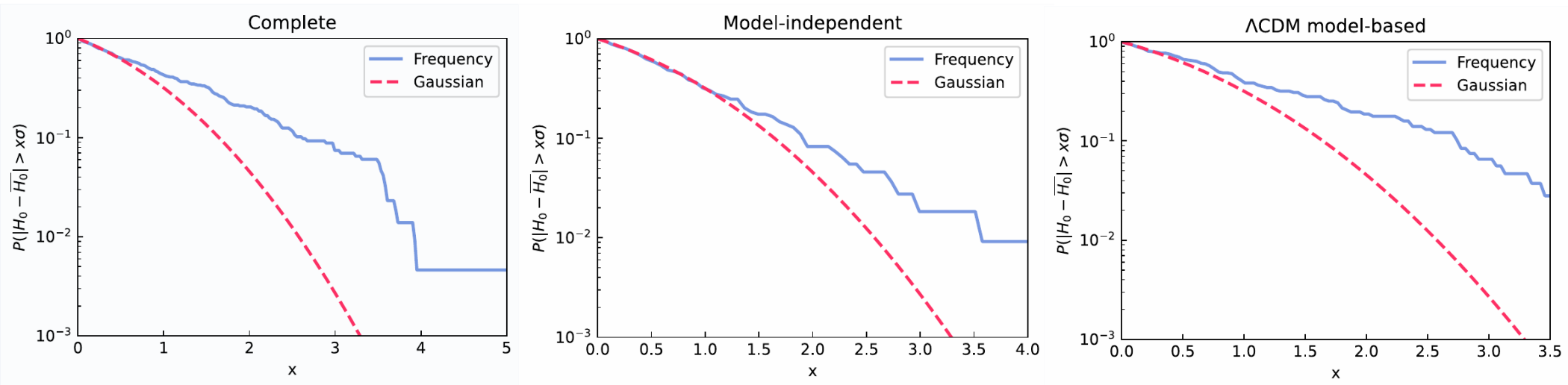
the frequency of deviations larger than  $x\sigma$  from  $\bar{H}_0$



measures the degree to which the sample deviates from the Gaussian

# 4. Estimating the real degree of tension

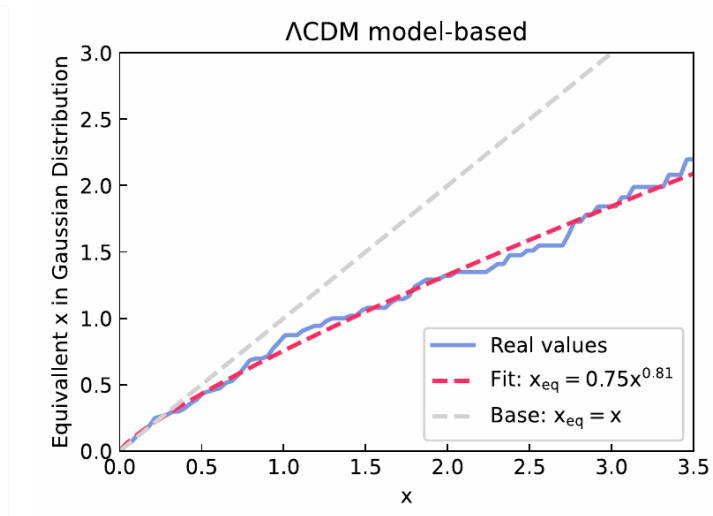
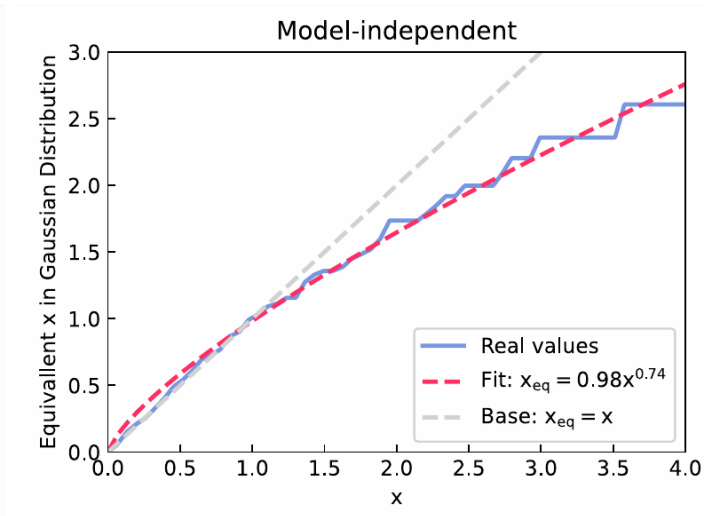
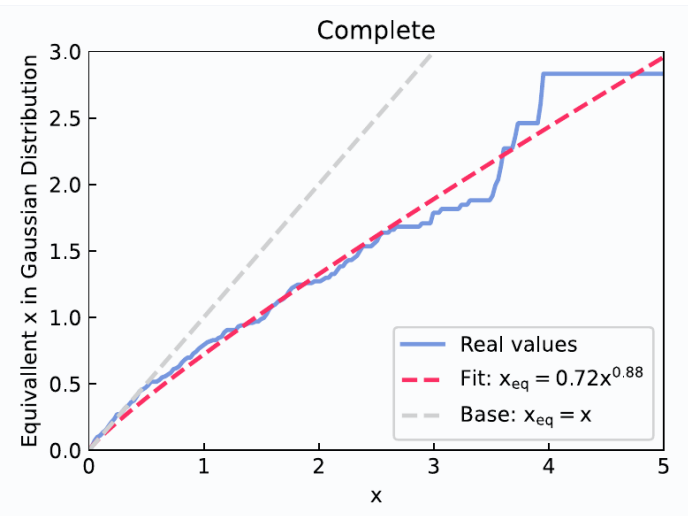
the frequency of deviations larger than  $x\sigma$  from  $\overline{H_0}$



$\implies$  define  $x_{eq}$  – the equivalent deviation between the probabilities of the the Gaussian distribution and the real frequency

# 4. Estimating the real degree of tension

fit a function  $x_{eq} = a * x^b$



i.e. when the deviation of  $H_{0,i}$  from  $\overline{H_0}$  is  $x\sigma$ , it is actually  $x_{eq}\sigma$

## 4. Estimating the real degree of tension

$$x_{\text{eq}} = (0.719 \pm 0.013)x^{0.879 \pm 0.014} \quad (\text{Complete})$$

$$x_{\text{eq}} = (0.983 \pm 0.012)x^{0.744 \pm 0.011} \quad (\text{Model-independent})$$

$$x_{\text{eq}} = (0.750 \pm 0.007)x^{0.818 \pm 0.009} \quad (\Lambda\text{CDM model-based})$$

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$5\sigma$  tension reported by Riess et al. (2022) becomes:

$$x_{\text{eq}} = 0.719 * x^{0.879} = 0.719 * 5^{0.879} = (2.96 \pm 0.12)\sigma \approx 3\sigma$$



## 4. Estimating the real degree of tension

$\chi_{eq}$  calibration should be independent of the data that are tested, so authors recalibrate  $\chi_{eq}$  with data from 1976-2019 and get:

$$\chi_{eq} = 0.83 * \chi^{0.62} = 0.83 * 5^{0.62} = 2.25\sigma$$

# 5. The effect of correlated data

Data divided into 5 samples after removing correlated datasets:

- **complete** (216 → 152)
- **model-independent** (109 → 71)
- **$\Lambda$ CDM model-based** (107 → 81)
- **$H_0 < 71 \text{ km s}^{-1} \text{ Mpc}^{-1}$**  (118 → 85)
- **$H_0 \geq 71 \text{ km s}^{-1} \text{ Mpc}^{-1}$**  (98 → 67)

# 5. The effect of correlated data

**Table 4.** The results of dip tests, weighted averages and  $\chi^2$  fittings in three categories after removing some of the correlated measurements. The  $Q$  values measure the statistical significance for  $\chi^2$  fittings, and the  $p$  values measure the significance of the unimodality for dip tests.

	Number	$\overline{H_0}$ ( $\sigma$ ) ( $\text{km s}^{-1} \text{Mpc}^{-1}$ )	$\chi^2$	$Q$	$p$
Complete	152	$69.25 \pm 0.13$	357.49	$8.68 \times 10^{-19}$	0.12
Model-independent	71	$70.32 \pm 0.25$	123.61	$8.20 \times 10^{-5}$	0.62
$\Lambda$ CDM model-based	81	$68.99 \pm 0.15$	193.89	$1.91 \times 10^{-11}$	0.13
$H_0 < 71 \text{ km s}^{-1} \text{Mpc}^{-1}$	85	$68.84 \pm 0.09$	81.13	0.57	0.69
$H_0 \geq 71 \text{ km s}^{-1} \text{Mpc}^{-1}$	67	$73.65 \pm 0.15$	18.35	$\sim 1.00$	0.98

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no more multimodality in the **complete** sample (before:  $p=0.01$ )

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$Q$  measures the probability that trends come from chance errors  
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# TAKEAWAY

- the Hartigans' **Dip Test** is good to test multimodality
- the **bimodal distribution of  $H_0$**  is also present in a model-independent sample and a  $\Lambda$ CDM model-based sample
- the deviation of  $H_0$  measurements with respect to  $\bar{H}_0$  are larger than expected from their error bars if they follow a Gaussian distribution ==>  **$5\sigma$  tension may in fact be a  $3\sigma$  tension**
- there is a **compatibility with a Gaussian** distribution for samples where of  $H_0 < 71$  km s<sup>-1</sup> Mpc<sup>-1</sup> and where  $H_0 \geq 71$  km s<sup>-1</sup> Mpc<sup>-1</sup> without the need to remove outliers ==> error underestimation related to methodology (systematics)